An Exponential Regression model for estimating daily milk yields

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Backgrounds

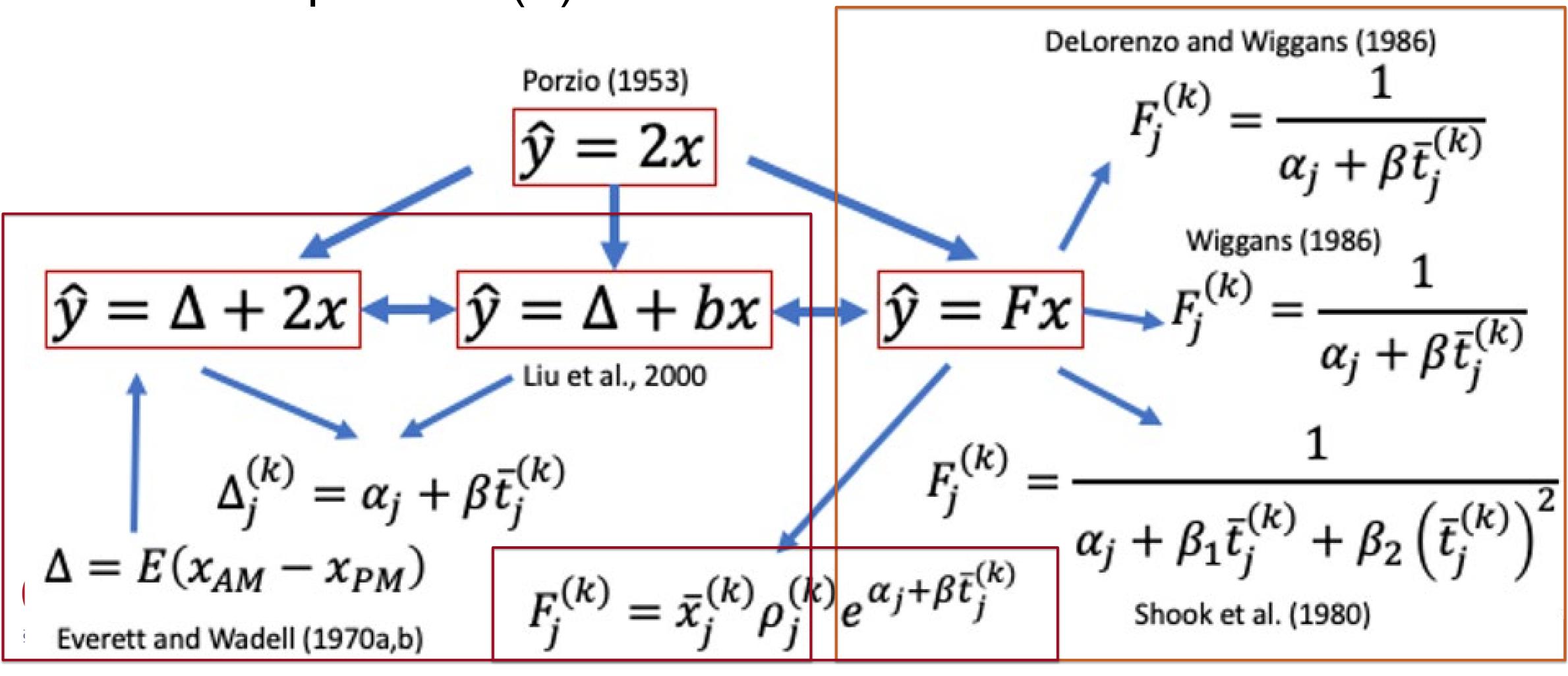
COUNCIL ON DAIRY CATTLE BREEDING

- A cow is typically milked two or more times each day during her lactation, and not all those milkings are weighed or sampled.
- The initial morning (AM)-evening (PM) milking plan alternately sampled AM or PM milking on test day throughout the lactation. Daily yield (milk, fat, and protein) was estimated by two times sampled AM or PM yield on each test day, assuming equal AM and PM milking intervals (Porzio, 1953). However, AM and PM milking intervals can vary, and milk secretion rates may be different between day and night Everrett and Wadell, 1970a, b).



Yield correction factors

and multiplicative (F) correction factors





\Box Various methods have been proposed focusing on additive (Δ)

Goals

QRe-evaluate the performance of existing statistical models, compared to the recently proposed exponential regression model, using crossvalidation.





Characterize ACF and MCF obtained from various

Additive correction factor (ACF) models

Regression with categorical variables (Everett and Wadell,

 $\mathcal{W}_{PM} = f(\theta) + \epsilon$ $\hat{y} = f(\hat{\theta}) + 2x_{PM}$

• A general form:

 $y = f(\theta) + bx + \epsilon$ dependent variables. COUNCIL ON DAIRY CATTLE BREEDIN



where $f(\theta)$ is a function with categorical or/and continuous

ACF evaluated for each MIC, say k

Regression with categorical variables, say MIC

$$\Delta_j^{(k)} = E(f(\theta|j, MIC) =$$

- where else = all categorical variables where applicable Regression model on discretized milking interval

$$\Delta_j^{(k)} = \hat{\alpha}_j + \hat{\beta} \bar{t}_j^{(k)}$$

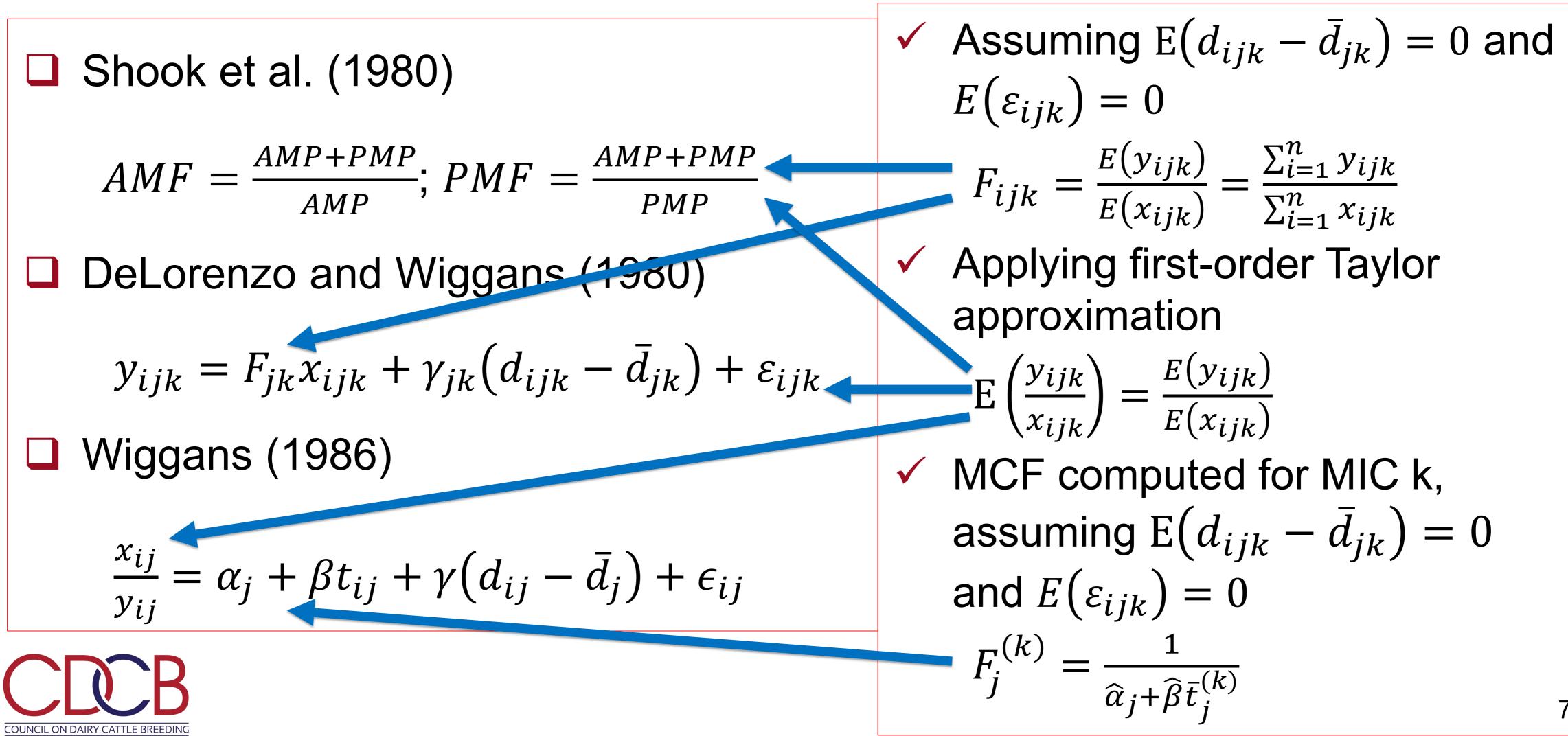




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k, else)

Multiplicative correction factor (MCF) models





Challenges: ACF models

- For regression with discrete MIC, the number of MIC coupled with other categorical variables increased substantially as more variables were considered.
 - two parities. Then, there would be $20 \times 4 \times 4 \times 4 \times 2 = 2,560$ specific classes for which ACF would need to be estimated while considering all these categorical variables at the same time.
- For example, consider twenty MIC, four herd location regions, four years, four seasons, and For regression with continuous milk interval, systematic biases arose from discretizing milking interval.

$$\Delta_{j}^{(k)} = E\left(\alpha_{j} + \beta t_{ij}^{(k)}\right) = \alpha_{j} + \beta E\left(\bar{t}_{j}^{(k)} + \left(t_{ij}^{(k)} - \bar{t}_{j}^{(k)}\right)\right) = \alpha_{j} + \beta \bar{t}_{j}^{(k)} + \beta E\left(t_{ij}^{(k)} - \bar{t}_{j}^{(k)}\right)$$
$$= \alpha_{j} + \beta \bar{t}_{j}^{(k)}$$

The above holds assuming $E\left(t_{ii}^{(k)} - \overline{t}_{i}^{(k)}\right)$ COUNCIL ON DAIRY CATTLE BREEDING



$${}_{j}^{(k)} = E\left(t_{ij}^{(k)}\right) - \overline{t}_{j}^{(k)} = 0.$$

Challenges: MCF models

An MCF model is challenged by the "ratio" problem because each has a ratio dependent variable in the data density or smoothing function. **Shook et al. (1980)**

 $\frac{PMP_{2k}}{AMP_{2k}+PMP_{2k}} = \alpha + \beta_1 t_k + \beta_2 t_k^2 + \epsilon_k;$

DeLorenzo and Wiggans

$$\frac{1}{F_j^{(k)}} = \alpha_j + \beta_j \overline{t}_{jk} + \epsilon_{jk}; F_j^{(k)} = \frac{1}{\widehat{\alpha}_j}$$

Wiggans (1986)

$$\frac{\alpha_{ij}}{\gamma_{ij}} = \alpha_j + \beta t_{ij} + \gamma (d_{ij} - \bar{d}_j) + \epsilon_{ij}$$

$$\begin{aligned} \hat{y}_{ij} + \gamma \left(d_{ij} - \bar{d}_j \right) + \epsilon_{ij}; \ F_j^{(k)} &= \frac{1}{\widehat{\alpha}_j + \widehat{\beta} \bar{t}_j^{(k)}}; \\ \hat{y}_{ij} &= x_{ij} * F_j^{(k)} + \widehat{\gamma} \left(d_{ij} - \bar{d}_j \right) \end{aligned}$$





$$AMF_{2k} = \frac{1}{\widehat{\alpha} + \widehat{\beta}_1 \overline{t}_k + \widehat{\beta}_2 \overline{t}_k^2}$$

$$\frac{1}{\hat{\beta}\bar{t}_{j}^{(k)}}$$

Challenges: The "ratio" problem

Bias due to dropping (or missing) main effect variables

 $\frac{y}{x} = \alpha \beta t + \epsilon$

 $y = \alpha x + \beta t x + \epsilon x$ \Box Bias due to additional measurement errors (δ) $\frac{y}{x+\delta} = + \beta t + \epsilon$

 $\frac{y}{x} = (\alpha + \beta t + \epsilon) + \left(\alpha \frac{\delta}{x} + \beta \frac{t\delta}{x}\right)$ COUNCIL ON DAIRY CATTLE BREEDING

$$+\epsilon_{ij}\frac{\delta}{x}$$



An exponential regression model

$$y_{ij} = x_{ij}^b e^{(\alpha_j + \beta t_j + \gamma (d_{ij} - \bar{d}_j) - \bar{d}_j)}$$

Logarithm transformation

$$log(y_{ij}) = \alpha_j + \beta t_j + \gamma (d_{ij} - \overline{d}_j) + blog(x_{ij}) + \epsilon_{ij}$$

nalogous to exponential growth function

$$y_{ij} \approx x_{ij}^b (1+1.718)^{\alpha_j + \beta t_j + \gamma \left(d_{ij} - \overline{d}_j\right) + \epsilon_{ij}}$$





□ Again, consider milking interval and days in milk (Wu et al., 2022) $+\epsilon_{ij}$

Estimating daily milk yield

Direct approach $\hat{y}_{ii} = \chi_{ii}^{\hat{b}} e^{\left(\hat{\alpha}_{j} + \hat{\beta}t_{j} + \hat{\gamma}(d_{ij} - \bar{d}_{j})\right)}$ Indirect approach (via MCF) $\hat{y}_{ij} = x_{ii}^{(k)} \times F_i^{(k)}$ $F_{j}^{(k)} = \frac{E(y_{ij}^{(k)})}{E(x_{ii}^{(k)})} = \rho_{j}^{(k)} \left(\bar{x}_{j}^{(k)}\right)^{b-1} e^{\alpha_{j} + \beta \bar{t}_{j}^{(k)} + \gamma \left(d_{jk} - \bar{d}_{j}^{(k)}\right)}$ where: $\rho_i^{(k)} = e^{\frac{1}{2} \left(V(y_{ij}^{(k)}) (\bar{y}_j^{(k)})^{-2} - bV(x_{ij}^{(k)}) (\bar{x}_j^{(k)})^{-2} \right)}$







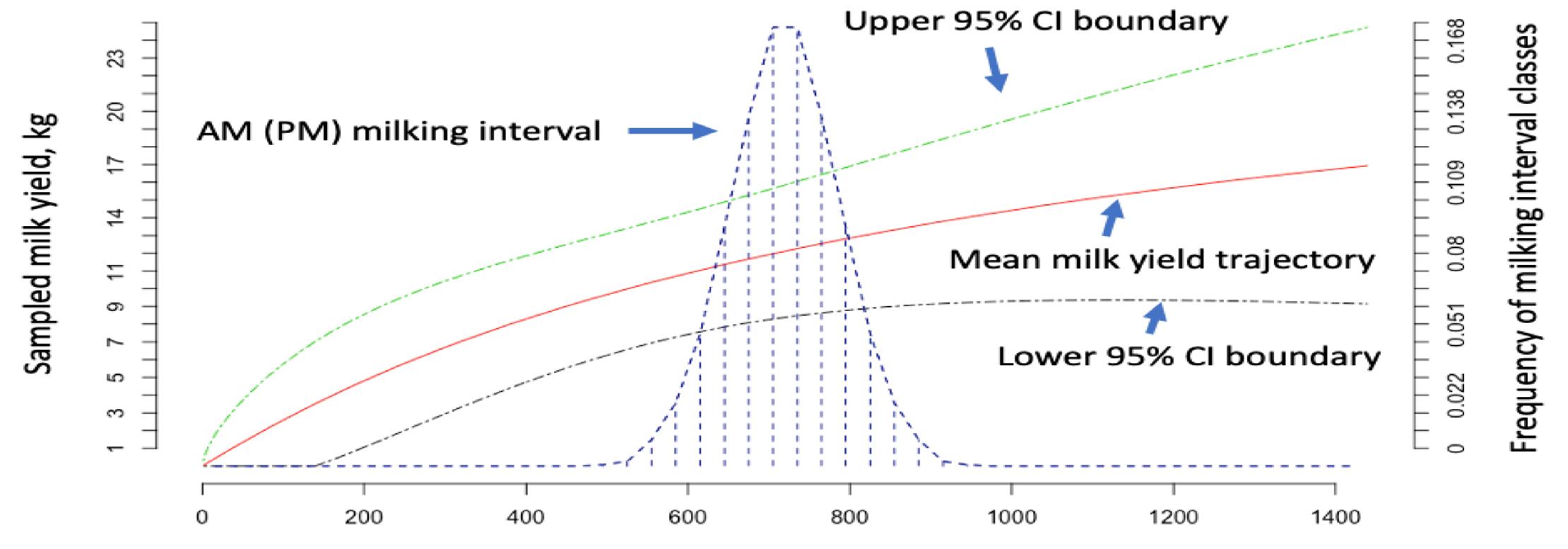
A simulation study

- □ Daily milk yields were simulated based on a modified Michael-Menten function (Klopcic et al., 2012), where $y_{720} \sim TN(12,2)$ and $k \sim TN(0.8,0.1)$.
- Thirty replicates were simulated; Each replicate consists of 3,000 cows; 2/3 for training and 1/3 for testing
- The performance (accuracy and decomposed MSE) for eight
 - methods were evaluated by 10-fold cross-validation.





Mean and 95% confidential intervals of daily milk yield and frequency distribution of AM (PM) milking intervals.



Minutes in milk





Decomposed MSE and Accuracy

Models	Variance	Bias ²	MSE	Accuracy
M 1	4.90E-04	0.486	0.486	0.968
M2A	7.40E-05	0.448	0.448	0.971
M2B	7.50E-05	0.480	0.480	0.968
M3A	7.40E-05	0.435	0.435	0.972
M3B	7.40E-05	0.465	0.465	0.970
M4	4.60E-04	0.422	0.422	0.972
M5	4.30E-04	0.420	0.421	0.972
M6A	5.90E-05	0.386	0.386	0.975
M6B	5.90E-05	0.417	0.417	0.973
M7A	5.90E-05	0.376	0.376	0.976
M7B	7.60E-05	0.385	0.385	0.975





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Comparing model parameters

Statistical	Model parameters					
models	α_{AM}	α_{PM}	β	b	Linear regression fit	Correlation
M2A	14.17	14.19	-1.182	2.0	AM: $y = 0.822 + 0.966 \hat{y}$	0.985
	(0.095)	(0.095)	(0.008)		PM: $y = 0.580 + 0.976 \hat{y}$	0.985
M3A	14.46	14.48	-1.147	1.942	AM: $y = 0.123 + 0.995 \hat{y}$	0.986
	(0.096)	(0.096)	(0.008)	(0.004)	PM: $y = -0.126 + 1.005 j$	0.985
M6A	0.208	0.208	0.024		AM: $y = 0.684 + 0.972 \hat{y}$	0.986
	(0.002)	(0.002)	(<0.001)		PM: $y = 0.577 + 0.976 \hat{y}$	0.986
M7A	1.324	1.324	-0.048	0.977	AM: $y = 0.102 + 0.996 \hat{y}$	0.988
	(0.005)	(0.005)	(<0.001)	(0.002)	PM: <i>y</i> = −0.009 + 1.001 j	0.987

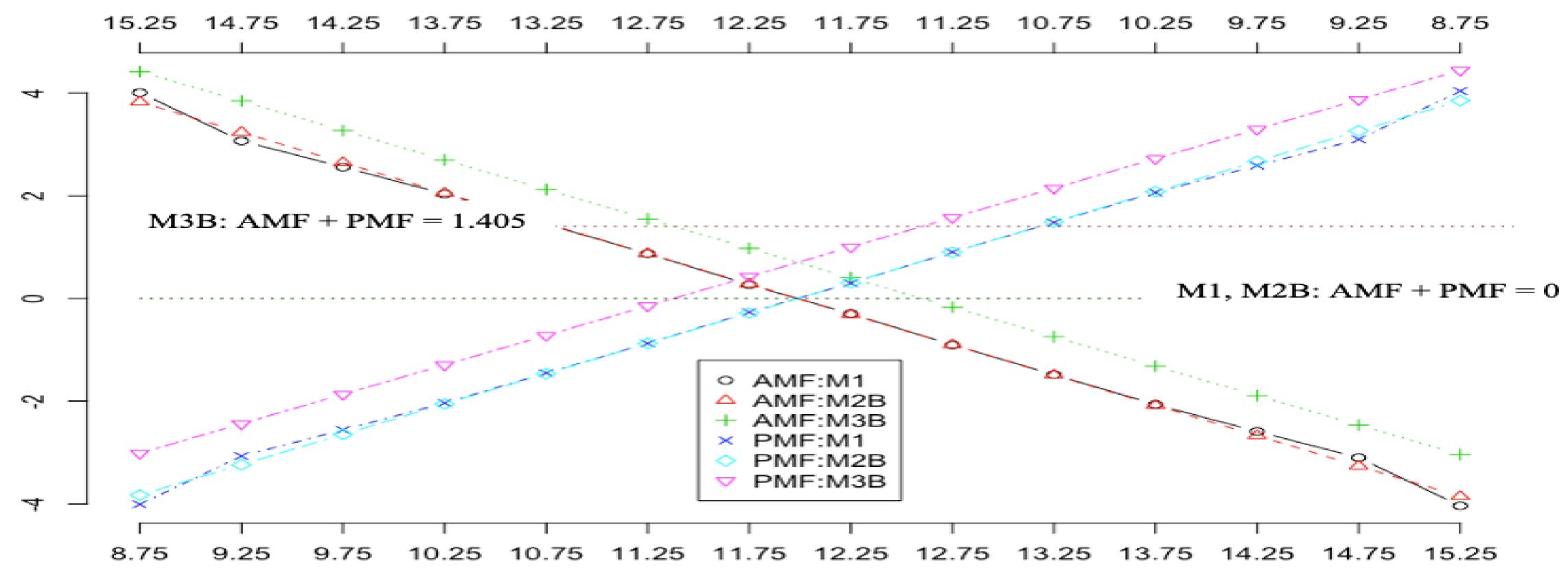








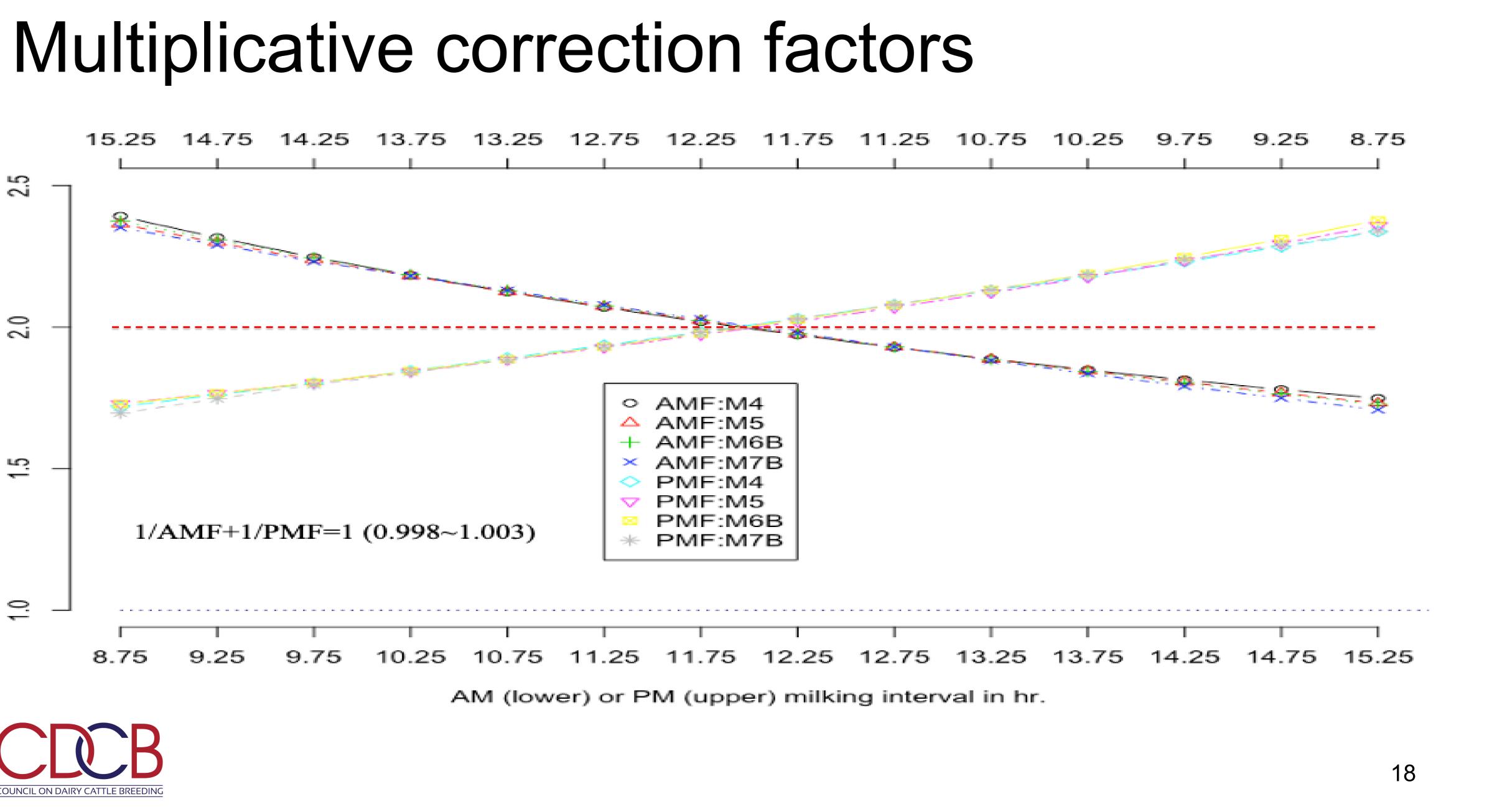
Additive correction factors







AM (lower) or PM (upper) milking interval in hr.





Conclusions

- Doubling AM (PM) milk yields was approximately taken with equal AM and PM milking intervals, but it was subject to large errors with unequal milking intervals. In contrast, the use of ACF and MCF has effectively reduced the biases due to unequal milking intervals.
- All the methods had high precision but they differ considerably in the accuracy of estimates. Comparably speaking, MCF and linear regression models had smaller MSE and higher accuracy than ACF.
- The exponential regression models had the smallest MSE and the greatest accuracies of all the model evaluated, thus providing a promising alternative tool for estimating DMY.

council on the estimation of daily fat and protein yields.



The methods were presented with the simulated AM and PM milking data, yet the principles are generally applicable to cows milked more than two times a day, and they 19

Comments and questions?



